Fine-grained Analysis for the Phase Transition of Moment Matching, with Application in Infinite-Armed Bandit Problem

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infinite bandits algorithm

- each arm's average reward is sampled from an unknown distribution, and each arm can be sampled further to obtain noisy estimates of the average reward of that arm
- We consider a general class of distribution functionals g(F). For the special case of median estimation, we identify a curious thresholding phenomenon

Goal

We aim to fine-grained analysis this thresholding behavior for median estimation.

Indicator Based functionals

Definition 1 (Indicator-based functionals). The functional g can be represented as

$$g(F) = \mathbb{E}[X \mid X \in S(F)]$$

for some set S(F), where $X \sim F$. The set S(F) is defined as follows:

$$S(F) = [F^{-1}(\alpha_1), F^{-1}(\alpha_2)], \quad 0 \le \alpha_1 \le \alpha_2 \le 1.$$

Functionals For Median

For median case, $\alpha_1 = \alpha_2 = \frac{1}{2}$

Median Estimation

Assumption1 There exist constants $c_2 > 0$ such that

▶
$$|F''(x)| \le c_2$$
 for $|x - \text{median}(F)| \lesssim \sqrt{\varepsilon}$.

Comment:

► The assumption precludes the distribution from being dumbbell-shaped.

Median Estimation

Le Cam's two-point lower bound

- ▶ Let F_1 and F_2 be two distributions with $|g(F_1) g(F_2)| \ge 2\varepsilon$,
- ► Le Cam's two-point lower bound gives:

$$\inf_{\hat{g}} \sup_{F \in \{F_1, F_2\}} \mathbb{P}_F\left(|\hat{g} - g(F)| \geq \varepsilon\right) \geq \frac{1}{4} \exp\left(-\underbrace{D_{\mathsf{KL}}(p_{\pi, F_1} \parallel p_{\pi, F_2})}_{\mathsf{Goal: Upper Bound}}\right).$$

where in the offline algorithm, $p_{\pi,F} = (F * \mathcal{N}(0,1/m))^{\otimes n}$.

Key quantity

Let \mathcal{F} denote the set of distributions satisfying Assumption 2 $\mathrm{KL}_{\sigma}(\varepsilon) \triangleq \min \Big\{ D_{\mathrm{KL}}(F_1 * \mathcal{N}(0, \sigma^2) \parallel F_2 * \mathcal{N}(0, \sigma^2)) : \\ F_1, F_2 \in \mathcal{F}, \ |F_1^{-1}(1/2) - F_2^{-1}(1/2)| \geq 2\varepsilon \Big\}.$

How to characterize $KL_{\sigma}(\varepsilon)$

▶ (Wang, Y., Baharav, T. Z., Han, Y., Jiao, J., Tse, D. (2022). Beyond the best: Estimating distribution functionals in infinite-armed bandits.) have proved: For $\varepsilon \in (0, 1/4)$, $\mathsf{KL}_{\sigma}(\varepsilon)$ can be characterized as follows:

$$\mathsf{KL}_{\sigma}(\varepsilon) \in \begin{cases} [C_1 \varepsilon^2, C_2 \varepsilon^2] & \text{if } \sigma \leq c \varepsilon^{1/2}, \\ \leq C(\theta, \kappa) \varepsilon^{\kappa} & \text{if } \sigma \geq \varepsilon^{1/2 - \theta}, \end{cases}$$

where $\theta \in (0, 1/4)$, $\kappa \in \mathbb{N}$ are arbitrary fixed parameters

Thresholding Phenomenon

$$\mathsf{KL}_{\sigma}(\varepsilon) \in \begin{cases} [C_1 \varepsilon^2, C_2 \varepsilon^2] & \text{if } \sigma \leq c \varepsilon^{1/2}, \\ \leq C(\theta, \kappa) \varepsilon^{\kappa} & \text{if } \sigma \geq \varepsilon^{1/2 - \theta}, \end{cases}$$

Intuition behind the thresholding phenomenon:

- When $\sigma = O(\varepsilon^{1/2})$, the "bandwidth" of $F_1 F_2$ exceeds that of $\mathcal{N}(0, \sigma^2)$, and the convolution is effectively using $\mathcal{N}(0, \sigma^2)$ as a Gaussian kernel (which preserves polynomials up to order 2) for smoothing $F_1 F_2$.
- When $\sigma \gg \varepsilon^{1/2}$, the "bandwidth" of $F_1 F_2$ could be smaller than $\mathcal{N}(0, \sigma^2)$, and the convolution is effectively using $F_1 F_2$ as a kernel(which could preserve polynomials up to any desired order) for smoothing $\mathcal{N}(0, \sigma^2)$.

Main Result

Finer characterization

In this project, we improved the result to be following

$$\textit{KL}_{\sigma}(\varepsilon) \asymp \begin{cases} \Theta(\varepsilon^2) & \text{if } \varepsilon \geq \sigma^2, \\ \varepsilon^{\kappa} \exp[-\Theta(\frac{\sigma^2}{\varepsilon})] & \text{if } \sigma^4 < \varepsilon < \sigma^2 \\ \exp[-\Theta\left(\frac{1}{\sqrt{\varepsilon}}\log(\frac{\sigma^4}{\varepsilon})\right)] & \text{if } \varepsilon < \sigma^4 \end{cases}$$

Note

For the second region, we have not achieved exact characterization.

Upper Bound

Characterize Upper bound

We could upper bound $KL_{\sigma}(\varepsilon)$, by the its χ^2 distance, similar strategy as (Yihong Wu and Pengkun Yang 2020):

$$KL(P,Q) \le \chi^2(P,Q) \lesssim \sum_{j\ge 1} \frac{(\Delta_j)^2}{j!\sigma^{2j}}$$

where $P = \mu * N(0, \sigma^2)$, $Q = \nu * N(0, \sigma^2)$, μ, ν are two probability measure supported on [-1, 1], and $N(0, \sigma^2)$ is a centered gaussian with variance σ^2 , Δ_j is the jth moment difference between P and Q

Proof.

Write densities of two mixture distributions $\nu * N(0, \sigma^2)$ and $\nu' * N(0, \sigma^2)$ are

$$f(x) = \int \phi(x - u) d\nu(u) = \phi(x) \sum_{j \ge 1} H_j(\frac{x}{\sigma}) \frac{m_j(\nu)}{j!},$$

$$g(x) = \int \phi(x - u) d\nu'(u) = \phi(x) \sum_{j \ge 1} H_j(\frac{x}{\sigma}) \frac{m_j(\nu')}{j!},$$

apply Jensen's Inequality and some simplification we have

$$g(x) \ge \phi(x) \exp(-\sigma^2/2).$$

$$\chi^{2}(\nu * N(0,\sigma) \| \nu' * N(0,\sigma)) = \int \left(\frac{f(x) - g(x)}{g(x)}\right)^{2} dx$$

$$\leq e^{\frac{\sigma^{2}}{2}} \mathbb{E}\left[\left(\sum_{j \geq 1} \frac{H_{j}(Z) \Delta m_{j}}{j! \sigma^{2k}}\right)^{2}\right] = e^{\frac{\sigma^{2}}{2}} \sum_{j \geq 1} \frac{(\Delta m_{j})^{2}}{j! \sigma^{2k}},$$

where Z(0,1) and the last step follows from the orthogonality of Hermite polynomials: $\mathbb{E}[H_i(Z)H_i(Z)] = j!\mathbf{1}_{\{i=i\}}$.

With previous upper bound model, we now can upper bound $KL_{\sigma}(\varepsilon)$ by following strategy:

$$\mathsf{KL}_{\sigma}(\varepsilon) \lesssim \sum_{j=1}^{\infty} \frac{(m_j(F_1) - m_j(F_2))^2}{\sigma^{2j} j!} \sim \frac{(m_k(F_1) - m_k(F_2))^2}{\sigma^{2k} k!}$$

where we set
$$m_1(F1) = m_1(F1)...m_{k-1}(F1) = m_{k-1}(F2)$$
, $supp(F1) = supp(F2) = [-1, 1]$

Optimization Step

$$Vk = \max_{f} \int_{0}^{1} f(x)dx$$

st:
$$\begin{cases} ||f'||_{\infty} \leq 1\\ \int_{-1}^{1} f(x)x^{i}dx = 0, \text{ for } i = 1...k. \end{cases}$$

Duality Step

By setting g = f', optimization problem now becomes:

$$Vk = \max_{g} \int_{-1}^{1} g(x)(x \vee 0) dx$$

st:
$$\begin{cases} ||g||_{\infty} \leq 1 \\ \int_{-1}^{1} g(x)x^{i} dx = 0, \text{ for } i = 1...k. \end{cases}$$

$$Vk = \max_{g} \inf_{a_{1}, a_{2}, \dots, a_{k}} \left[\int_{-1}^{1} g(x) \left(x \vee 0 - \sum_{i=1}^{k} a_{i} x^{i} \right) dx : \|g\|_{\infty} \leq 1, \int_{-1}^{1} g(x) x^{i} dx = 0 \right]$$

$$\leq \inf_{a_{1}, a_{2}, \dots, a_{k}} \max_{g} \left[\int_{-1}^{1} g(x) \left(x \vee 0 - \sum_{i=1}^{k} a_{i} x^{i} \right) dx, \|g\|_{\infty} \leq 1 \right]$$

$$\leq \inf_{a_{1}, a_{2}, \dots, a_{k}} \left\| x \vee 0 - \sum_{i=1}^{k} a_{i} x^{i} \right\|_{L^{1}[-1, 1]}.$$

L_1 estimation step

By approximation theory
$$\|x \vee 0 - \sum_{i=1}^k a_i x^i\|_{L^1[-1,1]} \begin{cases} \lesssim \frac{1}{k^{2-\delta}} \\ \gtrsim \frac{1}{k^2} \end{cases}$$

By median difference assumption, we required

 $t^2V_k=rac{t^2}{k^2}\sim V_k(t)\geq arepsilon$, so we have $t_{\min}=k\sqrt{arepsilon}$, then:

$$\mathsf{KL}_{\sigma}(\varepsilon) \preceq \frac{(m_k(F1) - m_k(F2))^2}{\sigma^{2k} k!} \leq \frac{(k\sqrt{\varepsilon})^2}{\sigma^{2k} k!} \asymp (\frac{k\varepsilon}{\sigma^2})^k$$

We know that $k^* \sim \frac{\sigma^2}{\varepsilon}$, $t^* = \frac{\sigma^2}{\sqrt{\varepsilon}}$, which divide this upper bound into two region for $t^* < 1$ or $t^* > 1$

Lower Bound

We now construct a strategy for lower bound $KL_{\sigma}(\varepsilon)$, by the its Hellinger distance.

$$\mathit{KL}(P,Q) \geq H^2(P,Q) \gtrsim \sum_{j=1}^{(rac{t}{\sigma})^2} 2^{-j} rac{\Delta_j^2}{\sigma^{2j} j! e^{rac{t\sqrt{j}}{\sigma}}} + \sum_{j > rac{t}{\sigma}^2} rac{\Delta_j^2}{t^{2j}}$$

Here is proof sketch:

Proof.

- ▶ Observe $KL(P,Q) \ge H^2(P,Q) \asymp \int (\sqrt{p} \sqrt{q})^2 \asymp \int \frac{(p-q)^2}{p+q} \ge \frac{(\int f(dp-dq))^2}{\int f^2(dp+dq)} = \sup_f \frac{(\mathbb{E}_p[f] \mathbb{E}_q[f])^2}{\mathbb{E}_p[f^2] + \mathcal{E}_q[f^2]}$
- Then we use Hermite polynomial approximate $(\mathbb{E}_p[f] \mathbb{E}_q[f])^2$ and $\mathbb{E}_p[f^2]$, $\mathbb{E}_q[f^2]$



Characterize Lower Bound

By Cauchy-Swartz inequality, $\exists \{w_j\}$, then the optimization problem becomes

$$\begin{split} Vk^* & \geq \max_{w_j} \left(\frac{1}{\sum_{j=0}^k w_j^2 j! \sigma^{2j} A_j(\sigma)} \right) \left(\int_{-t}^t \sum_{j=1}^k w_j f(x) x^j dx \right)^2 \\ & \geq \max_{w_j} \left(\frac{1}{\sum_{j=0}^k w_j^2 j! \sigma^{2j} A_j(\sigma)} \right) \left(\int_{-t}^t (x \vee 0) h(x) dx + (\sum_{j=1}^k w_j f(x) x^j dx - \int_{-t}^t (x \vee 0) h(x) dx) \right) \\ & \geq \max_{w_j} \left(\frac{1}{\sum_{j=0}^k w_j^2 j! \sigma^{2j} A_j(\sigma)} \right) \underbrace{\left(\varepsilon - \| \sum_j w_j x^{j+1} - \max(x, 0) \|_{L1[-t, t]} \right)^2}_{B} \end{split}$$

L_1 estimation

 L_1 estimation tells us $\|\sum_j w_j x^{j+1} - \max(x,0)\|_{L^1[-1,1]} \leq \frac{1}{k^{2-\delta}}$, and in order to make $B \approx \varepsilon^2$, we then require $\frac{t^2}{k^2} \approx \varepsilon$, which means $k \sim \frac{t}{\sqrt{\varepsilon}}$. Approximation theory itself gives us $w_i = O(t^{-j})$

Characterize Lower Bound

$$\varepsilon < \sigma^4$$

Then we left with the characterization of A, for $\varepsilon < \sigma^4$:

$$A = \left(\frac{1}{\sum_{j=0}^k w_j^2 j! \sigma^{2j} A_j(\sigma)}\right) \ge \frac{1}{\left(k\sigma^2 t^2\right)^k} \ge \exp\left[-\frac{1}{\sqrt{\varepsilon}} \log\left(\frac{\sigma^4}{\varepsilon}\right)\right]$$

Combined with upper bound, for this region, we have then proved

$$\mathit{KL}_{\sigma}(arepsilon) symp \exp[-\Theta\left(rac{1}{\sqrt{arepsilon}}\log(rac{\sigma^4}{arepsilon})
ight)]$$

$$\sigma^4 < \varepsilon < \sigma^2$$

We conjectured that $KL_{\sigma}(\varepsilon) \simeq \varepsilon^{\kappa} \exp[-\Theta(\frac{\sigma^{2}}{\varepsilon})]$

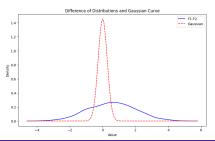
Future Work

Local Moment Matching

With small support t, the characterization is more delicate. We observe a following trade-off that $(t \sim \frac{\sigma^2}{\sqrt{\epsilon}})$

$$\mathit{KL}(1) \geq \max_t (\mathit{KL}(t) - \exp(-\frac{t^2}{\sigma^2}))$$

Goal: try to find the optimal t^* that balance this trade-off



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